

### Optimal Designs and Adaptive Randomization Techniques in Clinical Trials

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Optimal Designs & Adaptive Randomization



### Outline



- 2 Adaptive Designs
- 3 Adaptive Randomization
- Optimal Allocation
- 8 Randomization procedures used to target optimal allocation
  - Allocation-Adaptive Randomization (AAR)
  - Response-Adaptive Randomization (RAR)
  - Covariate-Adaptive Randomization (CAR)
  - Covariate-Adjusted Response-Adaptive Randomization (CARA)



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### **Clinical Trials**

Clinical trials are prospective biomedical or behavioural research studies on **human subjects** that are designed to answer specific questions about biomedical or behavioural interventions:

- novel vaccines,
- drugs,
- treatments,
- functional foods,
- dietary supplements,
- devices,
- new ways of using known interventions

generating safety and efficacy data.



### **Clinical Trials**

- Typically randomized, double-blind, placebo and/or active controlled study designs.
- The most common objective of a randomized clinical trial is to test the hypothesis that a new treatment is better than the standard fo care in the population with the disease.





Optimal Designs & Adaptive Randomization





### **Clinical Drug Development**



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### **Clinical Drug Development**



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## Very Expensive (!) Process

#### It costs $\sim$ **\$2.56B** to bring a New Medicine to a Market!!!<sup>1</sup>

March 10, 2016

# Tufts CSDD Assessment of Cost to Develop and Win Marketing Approval for a New Drug Now Published

BOSTON – March 10, 2016 – The most recent analysis by the Tufts Center for the Study of Drug Development of the average cost to develop and gain marketing approval for a new drug—pegged at \$2.558 billion—has been published in the *Journal of Health Economics*, it was announced today.

<sup>&</sup>lt;sup>1</sup><u>The Tufts Center for the Study of Drug Development</u> (independent, academic, non-profit research group at Tufts University in Boston, Massachusetts).



### Call for Innovation

Recognizing the challenges for modern drug development, the **FDA** released the **Critical Path Initiative** to encourage use of innovative tools to streamline drug development:

- Biomarkers
- Innovative trial designs
- Pharmacometrics
- Bioinformatics





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### Adaptive Designs: Motivation

The commonly used way of conducting a clinical trial:

- *n* subjects are involved in a study.
- $K \ge 2$  treatment arms are investigated (selected).
- Subjects are allocated to treatments according given proportions (*equal in many cases*).
- Given subjects' responses (*efficacy variables*), statistical inference on drug(s) properties is performed.



### Adaptive Designs: Motivation

- At the planning stage of a trial, various assumptions (treatment effect, variance, dropout rate) must be made. *Inaccurate assumptions increase risk of trial failure*.
- Having an option to modify trial design adaptively, based on interim results can help reduced uncertainty and improve decision-making.
- Adaptation is a design feature, not a remedy for poor planning



# Adaptive Designs: Motivation

#### Analogy Between Adaptive Designs and Swiss Army $Knife^2$

- (a) Simple scissor
  - Optimal tool for a specific task.



- (b) Swiss Army knife
  - Versatile tool that combines several individual functions in a single unit.



Giant Swiss Army knife

- Functions for every perceivable need.
- Looks impressive.
- Highly impracticable.
- Very expensive.



<sup>2</sup>Bretz F, Gallo P, Maurer W (2017) The Swiss Army knife among clinical trial designs? Clinical Trials 14(5), 417-424

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### Types of Adaptation Available

- Adaptive allocation rule change in the randomization procedure to modify the allocation proportion or the number of treatment arms
- Adaptive sampling rule change in the number of study subjects
- Adaptive stopping rule early stopping due to efficacy, futility, or safety
- Adaptive decision rule change in the way decisions will be made about the trial (e.g., change of endpoint, change of test statistics, etc.)

In practice, combinations of adaptive rules are used.

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Let us consider a clinical trial with n subjects involved.

Let

- $T_1, T_2, \ldots, T_n$ ,  $T_j = k$   $(j = 1, 2, \ldots, n; k = 1, 2, \ldots, K)$ be a sequence of treatment assignments;
- $X_1, X_2, ..., X_n$ ,  $X_j = x_k$  (j = 1, 2, ..., n; k = 1, 2, ..., K)be a sequence of responses;
- $Z_1, Z_2, \dots, Z_n, \quad Z_j = \left(z_j^{(1)}, z_j^{(2)}, \dots, z_j^{(r)}\right)' \quad (j = 1, 2, \dots, n)$ be a sequence of subjects' covariates;



#### • Allocation-Adaptive Randomization (AAR)

 $P_k(j) = \Pr\left(T_j = k | T_1, \dots, T_{j-1}\right).$ 

• Response-Adaptive Randomization (RAR)

 $P_k(j) = \Pr(T_j = k | T_1, \dots, T_{j-1}, X_1, \dots, X_{j-1})$ 

• Covariate-Adaptive Randomization (CAR)

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### Three Steps to Develop an Optimal Randomization Procedure

- Derive an *optimal allocation* to satisfy selected experimental objectives.
- Choose *randomization procedure(s)* to implement the desired optimal allocation.
- Evaluate *operating characteristics* of the optimal randomization procedures under a variety of standard to worst-case scenarios

Select one that has best performance for use in practice



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Consider a clinical trial with  $K \ge 2$  treatment arms for which we want to find an *optimal design*. Then, one has to perform the following steps:

- Study objectives are formulated as a mathematical problem (usually as an optimization problem).
- **Optimal allocation**  $\rho^* = (\rho_1^*, \rho_2^*, \dots, \rho_K^*)$  is found as a solution of the problem under constraints:

• 
$$0 \le \rho_k^* \le 1, \quad k = 1, 2, \dots, K.$$

• 
$$\sum_{k=1}^{K} \rho_k^* = 1.$$

③ A *randomization procedure* has to be constructed:

- it sequentially allocates subjects to treatments.
- the allocation proportion vector

#### $(n_1/n, n_2/n, \dots, n_K/n) \approx (\rho_1^*, \rho_2^*, \dots, \rho_K^*)$

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Optimal Designs & Adaptive Randomization



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- Two treatment groups: 1 and 2.
- $n = n_1 + n_2$  total sample size (fixed).
- $\rho^* \in (0,1)$  optimal allocation proportion for treatment 1 (to be determined) such that  $n_1 = n\rho^*$  and  $n_2 = n(1 \rho^*)$ .
- $Y_{jk} \sim Normal(\mu_k, \sigma^2)$  response of the *j*th (j = 1, 2, ..., n) patient in group k (k = 1, 2).
- **Objective**: Maximize power of *t*-test for testing  $H_0: \mu_1 = \mu_2$ .
- **Solution**:  $\rho^* = 0.5$  (equal number of patients should be assigned to treatments 1 and 2).



- Two treatment groups: 1 and 2.
- $n = n_1 + n_2$  total sample size (fixed).
- $Y_{jk} \sim Bernoulli(p_k)$  binary response of the *j*th (j = 1, 2, ..., n) patient in group k (k = 1, 2). Here,  $p_k = \Pr(Y_{jk} = 1)$  probability of success.
- Objective 1: Maximize power of Z-test for testing H<sub>0</sub>: p<sub>1</sub> = p<sub>2</sub>.
   Solution 1: ρ<sup>\*</sup> = √(p<sub>1</sub>q<sub>1</sub>)/√(p<sub>1</sub>q<sub>1</sub>) + √(p<sub>2</sub>q<sub>2</sub>).
- **Objective 2**: Minimize expected number of treatment failures subject to fixed power of Z-test under  $H_0: p_1 \neq p_2$ .

• Solution 2: 
$$\rho^* = \frac{\sqrt{p_1}}{\sqrt{p_1} + \sqrt{p_2}}$$
.



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D-optimal design for dose-finding studies with TTE outcomes

- $T \sim Weibull(\lambda, k), \quad f_T(t) = \frac{k}{\lambda} \left(\frac{t}{\lambda}\right)^{k-1} \exp\left(-\left(\frac{t}{\lambda}\right)^k\right)$
- Accelerated Failure Time (AFT) model:

$$\log (T) = \beta_0 + \beta_1 x + \beta_2 x^2 + b\varepsilon$$
$$\lambda = \exp (\beta_0 + \beta_1 x + \beta_2 x^2), \quad b = k^{-1}$$
$$\varepsilon \sim f_{\varepsilon}(u) = \exp (-\exp (-u)) - \underline{\text{extreme value distribution}}$$
$$x \in \mathcal{X} = [0, 1] \text{ is a treatment dose}$$

• The aim is to define a dose-response curve  $(Median \ TTE)$ 

$$Median(T|x) = \exp\left(\beta_0 + \beta_1 x + \beta_2 x^2\right)\log^b(2)$$

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*Optimal design vs. censoring time*:  $(\beta_0 = 1.9, \beta_1 = 0.6, \beta_2 = 2.8, b = 0.57721)$ 





**Dose-response fit, response rate = 50%:**  $(\beta_0 = 1.9, \beta_1 = 0.6, \beta_2 = 2.8, b = \gamma)$ 





### **Optimal Allocation: Summary**

- Choice of the target allocation ratio stems from the study objectives.
- Unequal allocation designs may be preferred over equal allocation designs for statistical (e.g. power/efficiency) and ethical reasons.
- Practical implementation of the chosen target allocation requires a *careful choice of randomization* (which can promote selected study objectives while *protecting study from experimental bias*)



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#### The following randomization procedures target allocation ratio

$$w = (w_1 : w_2 : \ldots : w_K), \quad w_k \in \mathbb{N}, \quad GCD(w_1, w_2, \ldots, w_K) = 1,$$
  
$$\boldsymbol{\rho}^* = (\rho_1, \rho_2, \ldots, \rho_K), \quad \rho_k = \frac{w_k}{\sum_{i=1}^K w_k}.$$

- Permuted Block Design (PBD)
- Block Urn Design (BUD)
- Mass Weighted Urn Design (MWUD)
- Drop-the-Loser Urn Design (DLUD)
- Doubly-Adaptive Biased Coin Design (DBCD)
- Maximum Entropy Constraint Balance Randomization (MaxEnt)



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# All the designs depend on a tweak parameter the choice of which is an open question!

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#### Charachteristics of randomization procedure:

• Imbalance

$$Imb(n) = \frac{1}{n} \sum_{j=1}^{n} \sqrt{\sum_{k=1}^{K} (N_k(j) - j\rho_k)^2}$$

• Forcing Index

$$FI(n) = \frac{1}{n} \sum_{j=1}^{n} \left( \sum_{k=1}^{K} (P_k(j) - \rho_k)^2 \right)$$



AAR

#### Imbalance plot





AAR

#### Forcing index plot





#### Imbalance vs. Forcing index





#### Distributions of final allocation proportions





#### Allocation ratio preserving (ARP) property





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Let us consider a *dose-finding trial with time-to-event outcomes* and *n* subjects involved.

$$Median(T|x) = \exp\left(\beta_0 + \beta_1 x + \beta_2 x^2\right)\log^b(2)$$

A D-optimal optimal allocation as well as optimal treatments (doses) are found as a solution of the following optimization problem

$$(\boldsymbol{x}^*, \boldsymbol{\rho}^*) = \arg \max_{(\boldsymbol{x}, \boldsymbol{\rho})} \det |\boldsymbol{FIM}(\boldsymbol{x}, \boldsymbol{\rho}, \boldsymbol{\theta})|, \quad \boldsymbol{\theta} = (\beta_0, \beta_1, \beta_2, b)$$

$$x_k \in [d_{min}, d_{max}], \quad k = 1, 2 \dots, K$$
s.t. 
$$\rho_k \in [0, 1], \quad k = 1, 2 \dots, K$$

$$\sum_{k=1}^{K} \rho_k = 1$$

Optimal Designs & Adaptive Randomization



#### "Ideal" case

- All subjects are available.
- They are just splitted into two groups

$$n_1 \approx n\rho_1^*, \quad n_2 \approx n\rho_2^*, \quad n_3 \approx n\rho_3^*.$$

(we assume that  $n_1$ ,  $n_2$  and  $n_3$  are rounded up to integers, if necessary)



However, there are potential problems with this approach:

- All n subjects are unavailable at the beginning of the trial.
- Model parameters θ = (β<sub>0</sub>, β<sub>1</sub>, β<sub>2</sub>, b) are unknown at the beginning of the trial. Basically, the trial itself is conducted to estimate θ (and perform other statistical analysis, such as, for example, hypothesis testing of H<sub>0</sub> : β<sub>1</sub> = β<sub>2</sub> = 0)
- Subjects cannot be just "split" into groups. They must be randomized in order to avoid selection bias.

**Randomization** is an essential component of any comparative experiment.



- In practice, the *n* subjects enter the trial sequentially and must be immediately randomized to one of K (= 3) treatments.
- The enrolled subjects generate data, and these data can be used to sequentially estimate the model parameters  $\boldsymbol{\theta} = (\beta_0, \beta_1, \beta_2, b)$ .
- Thus, we construct *response-adaptive randomization* procedures which converges to the desired optimal allocations.



#### Dose-finding studies for TTE outcomes



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- In most randomized phase III clinical trials patients are heterogeneous.
- Linear regression model with constant variance for the primary outcome the most efficient allocation is one for which treatment numbers are balanced, both overall and across selected covariates.
- Heteroscedastic or nonlinear model for the primary outcome
  - the concept of balance is different from the linear model case.
    - Covariate-balanced response-adaptive design.



Why is it important to force balance over known covariates? Let us consider a model where responses follow a linear regression model with constant varince

$$\boldsymbol{Y} = \boldsymbol{\mu} + \alpha \boldsymbol{t} + \beta \boldsymbol{z}_1 + \boldsymbol{\varepsilon},$$

where

- $\mu$  is the overall mean.
- $\alpha$  is the treatment effect.
- $\beta$  is the covariate effect.
- t is an *n*-vector of treatment assignments (whose elements are 1 or -1).
- $z_1$  is an *n*-vector of covariate values.

•  $\varepsilon \sim N(0, \sigma^2 I)$ .



Why is it important to force balance over known covariates? Let

- $\boldsymbol{\theta} = (\mu, \alpha, \beta)$  be a vector of unknown parameters;
- $X = [\mathbf{1}, \mathbf{t}, z_1]$  be a design matrix (for simplicity, assume that  $z_1$  is centred and scaled, that is,  $\mathbf{1}z'_1 = 0$ , and  $z_1z'_1 = 1$ );
- $\widehat{\boldsymbol{\theta}} = (\boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\boldsymbol{Y}$ . be the least square estimator of  $\boldsymbol{\theta}$ ; The variance-covariance matrix of  $\widehat{\boldsymbol{\theta}}$  is given by

$$\operatorname{Var}\left[\widehat{\boldsymbol{\theta}}\right] = \sigma^2 (\boldsymbol{X}' \boldsymbol{X})^{-1}.$$



Why is it important to force balance over known covariates? Particularly,

$$\operatorname{Var}\left[\widehat{\alpha}\right] = \frac{\sigma^2}{n - (\boldsymbol{z}'\boldsymbol{t})^2 - (\mathbf{1}'\boldsymbol{t})^2/n}.$$

It is minimized when both

- z't = 0 (balance over covariate).
- $\mathbf{1}'t = 0$  (balanced treatment assignments).



#### CAR procedures

- Pocock-Simon's "minimization" procedure (1975).
- Atkinson's biased coin design (1982).
- Covariate-balanced randomization design (Yuan and Huang, 2010)



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Two main reasons to consider **CARA**:

- *Statistical*: For nonlinear and heteroscedastic models, optimal allocation may not be balanced across treatment arms.
- **Ethical**: The degree and direction of treatment effect may differ for patient subgroups within a treatment  $\Rightarrow$  increase probability of assigning the treatment that is most efficacious given the patient's covariate profile (**personalized treatment**).



#### CARA Randomization for Survival Trials<sup>3</sup>

- *n* patients are enrolled sequentially and are randomized to treatments 1 and 2.
- For the *j*-th patient, survival time  $T_{jk}$ , conditional on covariates  $z_i$  is exponential with mean

$$\lambda_k(\boldsymbol{z}_j) = \exp\left(\boldsymbol{\theta}'_k \boldsymbol{z}_j\right), \quad k = 1, 2,$$

where  $\theta_k = (\theta_{k0}, \theta_{k1}, \dots, \theta_{kp})'$  and  $\boldsymbol{z}_j = (1, z_{1j}, \dots, z_{pj})'$ .

T<sub>jk</sub> is subject to independent right-censoring with τ<sub>j</sub> > 0.
t<sub>jk</sub> = min(T<sub>jk</sub>, τ<sub>j</sub>),
δ<sub>jk</sub> = 1<sub>t<sub>jk</sub>=T<sub>jk</sub>,
(t<sub>jk</sub>, δ<sub>jk</sub>) are independent, j = 1,..., n<sub>k</sub>, k = 1,2
</sub>

<sup>3</sup>Sverdlov O., Rosenberger W.F., Ryeznik Y. (2013) "Utility of covariate-adjusted response-adaptive randomization in survival trials" *Statistics in Biopharmaceutical Research* 5(1), p. 38-53.



Example: Redesigning a Survival Trial<sup>3</sup> (cetuximab trial in advanced colorectal cancer)

- In a 21-month period, n = 572 eligible patients were randomized at a 1 : 1 ratio among TRT 1 (cetuximab plus best supportive care) and TRT 2 (best supportive care alone).
- The primary endpoint was overall survival (OS).
- Effectiveness of cetuximab was significantly associated with *K-ras mutation status*:
  - Patients with *wild-type K-ras tumors* benefited from cetuximab (median OS, 9.5 vs. 4.8 months; HR for death, 0.55).
  - Patients with a *colorectal tumor bearing mutated K-ras* did not benefit from cetuximab (median OS, 4.6 vs. 4.5 months; HR for death, 0.98).

<sup>&</sup>lt;sup>3</sup>Karapetis, C. S., Khambata-Ford, S., Jonker, D. J., et al. (2008). "K-ras mutations and benefit from cetuximab in advanced colorectal cancer". The New England Journal of Medicine, 359, p. 17571765



#### Simulation Study Results (10,000 simulation runs)

<i>n</i> = 572	Pocock-Simon	CARA	RAR
$N_{\rm A}/n$ (S.D.)	0.500 (0.002)	0.588 (0.037)	0.583 (0.039)
$N_{A0} N_{B0}(S.D.)$	169 169(1)	211 127 (16)	197 141(14)
$N_{A1}   N_{B1} (S.D.)$	117 117(1)	125 109 (13)	137 97 (11)
Deaths (S.D.)	372 (11)	362 (12)	<b>366</b> (12)
Total Time (S.D.)	3076 (106)	<b>3155</b> (113)	<b>3132</b> (112)
$\hat{\theta}_{A0}$ (S.D.)	2.62 (0.11)	2.62 (0.10)	2.62 (0.10)
$\hat{\theta}_{A1}$ (S.D.)	-0.68 (0.16)	-0.68 (0.15)	-0.68 (0.15)
$\hat{\theta}_{B0}$ (S.D.)	1.87 (0.09)	1.86 (0.11)	1.87 (0.10)
$\hat{\theta}_{B1}$ (S.D.)	0.02 (0.14)	0.03 (0.16)	0.02 (0.16)

- Because of treatment-covariate interaction, CARA resulted in greater skewing to A in the wild-type K-ras subgroup than in the mutated K-ras subgroup, whereas RAR had similar degree of skewing in the subgroups.
- CARA and RAR had, on average, 10 and 6 fewer deaths and greater total survival time than the Pocock-Simon design.
- All three procedures had the same power and very similar M.L.E.'s.



#### Books

#### Monographs on Statistics and Applied Probability 130

Randomised Response-Adaptive Designs in Clinical Trials



Anthony C. Atkinson Atanu Biswas

CRC Press Solar Libert Longe

#### WILEY

The Theory of Response-Adaptive Randomization in Clinical Trials

fang Hu and William F. Rovenberger



Randomization in Clinical Trials Theory and Practice Second Edition

Wiley Series in Probability and Statistics

The second

William F. Rosenberger + John M. Lachin

WILEY

Chapman & Hall/CRC Biostatistics Series

#### Modern Adaptive Randomized Clinical Trials

Statistical and Practical Aspects



Edited by Oleksandr Sverdlov





# CIM Course for Master/PhD students (Period IV 2018)

- Introduction to clinical trials, adaptive designs, optimal adaptive randomization procedures.
- **2** Choice of a target allocation for a multi-arm clinical trial.
- **③** Adaptive randomization designs to implement optimal allocation:
  - Allocation-adaptive randomization
  - Covariate-adaptive randomization
  - Response-adaptive randomization
  - Covariate-adjusted response-adaptive randomization
- Case studies (with examples in R) and practical aspects of implementing adaptive randomization